

How Reliable is Your Regression Model's Uncertainty Under Real-World Distribution Shifts?

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October 31, 2023



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Supervisors: Thomas B. Schön (Uppsala) and Martin Danelljan (ETH Zürich).

I will defend my thesis Towards Accurate and Reliable Deep Regression Models on Nov 30.

My research focuses on probabilistic deep learning, and often includes regression problems, uncertainty estimation methods or energy-based models.



This presentation is mainly based on our recent TMLR paper:

How Reliable is Your Regression Model's Uncertainty Under Real-World Distribution Shifts? Fredrik K. Gustafsson, Martin Danelljan, Thomas B. Schön Transactions on Machine Learning Research (TMLR), 2023

Quite large parts of our previous CVPR Workshops paper will however also be covered, as this is highly relevant background material:

Evaluating Scalable Bayesian Deep Learning Methods for Robust Computer Vision *Fredrik K. Gustafsson, Martin Danelljan, Thomas B. Schön* The Conference on Computer Vision and Pattern Recognition Workshops (CVPR Workshops), 2020



How Reliable is Your Regression Model's Uncertainty Under Real-World Distribution Shifts? Fredrik K. Gustafsson, Martin Danelljan, Thomas B. Schön Transactions on Machine Learning Research (TMLR), 2023

- We propose a benchmark for testing the reliability of regression uncertainty estimation methods under real-world distribution shifts.
- We then employ our benchmark to evaluate many of the most common uncertainty estimation methods, as well as two state-of-the-art uncertainty scores from out-of-distribution detection.
- We find that while all methods are well calibrated when there is no distribution shift, they become highly overconfident on many of the benchmark datasets thus **uncovering important limitations** of current methods.
- This demonstrates that more work is required in order to develop truly reliable uncertainty estimation methods for regression.

Outline



- 1. Background: General Setting
- 2. Background: Predictive Uncertainty Estimation using Bayesian Deep Learning
- 3. Background: Prediction Intervals, Coverage & Calibration
- 4. Background: Selective Prediction
- 5. Summary of Contributions
- 6. Motivating Example
- 7. Proposed Benchmark
- 8. Evaluated Methods
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- 10. Main Actionable Takeaways

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In a **supervised regression problem**, the task is to predict a *continuous* target value $y^* \in \mathcal{Y} = \mathbb{R}^K$ for any given input $x^* \in \mathcal{X}$. To solve this, we are also given a training set of i.i.d. input-target pairs, $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^N, (x_i, y_i) \sim p(x, y)$.

In this presentation, the focus will be on the 1D case, i.e. when $\mathcal{Y} = \mathbb{R}$.

The input space \mathcal{X} will correspond to the space of images.



We view a **Deep Neural Network (DNN)** simply as a function $f_{\theta} : \mathcal{X} \to \mathcal{O}$, parameterized by $\theta \in \mathbb{R}^{P}$. This function maps inputs $x \in \mathcal{X}$ to outputs $f_{\theta}(x) \in \mathcal{O}$ in some output space \mathcal{O} .

We also divide the DNN f_{θ} into a *backbone feature extractor*, and one or more smaller *network heads*. The feature extractor takes x as input and outputs a feature vector g(x), which is then fed into the network heads, producing the final output $f_{\theta}(x) \in \mathcal{O}$.





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The most common and straightforward deep regression approach is to let the DNN f_{θ} directly output predicted targets, $\hat{y}(x) = f_{\theta}(x)$, training the DNN by minimizing e.g. the L2 loss over the training data, $J(\theta) = \sum_{i=1}^{N} (y_i - f_{\theta}(x_i))^2$.

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DNNs $f_{\theta} : \mathcal{X} \to \mathcal{O}$ have become the go-to approach within computer vision and many other domains due to their impressive predictive power. However, they generally fail to properly capture the uncertainty inherent in their predictions.

The approach of *Bayesian deep learning* aims to address this issue in a principled manner. It deals with predictive uncertainty by decomposing it into the distinct types of *aleatoric* and *epistemic* uncertainty.

Aleatoric uncertainty captures inherent and irreducible ambiguity in the data.

Epistemic uncertainty accounts for uncertainty in the DNN model parameters θ .



Given an input x, it is not always obvious what the correct target value y should be.

For example, what is the correct classification target for an image that contains both a cat and dog? Or, how about images with very low brightness, in which it is difficult to recognize any objects at all?

Aleatoric uncertainty captures this type of *inherent* and *irreducible* ambiguity that can be present in the inputs *x*.

Input-dependent aleatoric uncertainty arises whenever the target y is expected to be inherently more uncertain for some inputs x than others.

2.1 Aleatoric Uncertainty



This is true e.g. in semantic segmentation, where image pixels right at object boundaries are inherently more difficult to classify than pixels in the middle of objects.



2.1 Aleatoric Uncertainty



This is true also in automotive 3D object detection, where it is inherently more difficult to estimate the 3D position and size of distant or partially occluded vehicles.





To estimate input-dependent aleatoric uncertainty, the DNN $f_{\theta} : \mathcal{X} \to \mathcal{O}$ can be used to specify a model $p(y|x; \theta)$ of the conditional target distribution.

For example if a Gaussian model is used, $p(y|x; \theta) = \mathcal{N}(y; \mu_{\theta}(x), \sigma_{\theta}^2(x))$, the DNN outputs both a mean $\mu_{\theta}(x)$ and variance $\sigma_{\theta}^2(x)$ for each input x.

The mean can be taken as a prediction, $\hat{y}(x) = \mu_{\theta}(x)$, whereas the variance $\sigma_{\theta}^2(x)$ naturally can be interpreted as a measure of aleatoric uncertainty for this prediction.







The DNN f_{θ} can be trained by minimizing the negative log-likelihood (NLL) $\mathcal{L}(\theta)$,

$$\mathcal{L}(\theta) = \sum_{i=1}^{N} -\log p(y_i|x_i; \theta).$$

For the Gaussian model $p(y|x;\theta) = \mathcal{N}(y;\mu_{\theta}(x), \sigma_{\theta}^2(x))$, minimizing the NLL is equivalent to minimizing the following loss $J(\theta)$,

$$J(heta) = \sum_{i=1}^{N} rac{\left(y_i - \mu_{ heta}(x_i)
ight)^2}{\sigma_{ heta}^2(x_i)} + \log \sigma_{ heta}^2(x_i).$$



Using DNNs to specify models $p(y|x; \theta)$ of the conditional target distribution does however not capture **epistemic** uncertainty, as information about the uncertainty in the model parameters θ is disregarded.

Large epistemic uncertainty is present whenever a large set of model parameters explains the given training data (approximately) equally well.



This is often the case for DNNs, since the corresponding optimization landscapes are highly multi-modal.

Disregarding the epistemic model uncertainty can lead to highly confident yet incorrect predictions, especially for inputs x which are not well-represented by the training data.



Epistemic uncertainty can be estimated in a principled manner by performing *approximate Bayesian inference*.

Instead of just finding a single point estimate $\hat{\theta}$ of the model parameters θ , by minimizing the negative log-likelihood $\mathcal{L}(\theta) = \sum_{i=1}^{N} -\log p(y_i|x_i;\theta)$ over the training set $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^{N}$, Bayesian inference entails estimating the full *posterior* distribution $p(\theta|\mathcal{D})$.

The posterior $p(\theta|D)$ is obtained from the data likelihood $\prod_{i=1}^{N} p(y_i|x_i;\theta)$ and a chosen prior $p(\theta)$ by applying Bayes' theorem, $p(\theta|D) \propto \prod_{i=1}^{N} p(y_i|x_i;\theta)p(\theta)$.



The posterior $p(\theta|D) \propto \prod_{i=1}^{N} p(y_i|x_i; \theta) p(\theta)$ is then utilized to obtain the *predictive* posterior distribution p(y|x, D),

$$p(y|x, \mathcal{D}) = \int p(y|x; \theta) p(\theta|\mathcal{D}) d\theta$$

$$\approx \frac{1}{M} \sum_{m=1}^{M} p(y|x; \theta^{(m)}), \quad \theta^{(m)} \sim p(\theta|\mathcal{D}),$$
(1)

which captures both aleatoric and epistemic uncertainty.

In practice, obtaining samples from the true posterior $p(\theta|D)$ is virtually impossible for DNNs, requiring an approximate posterior $q(\theta) \approx p(\theta|D)$ to be used instead.

2.3 Illustrative Example





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A DNN f_{θ} trained to directly output predicted targets, $\hat{y}(x) = f_{\theta}(x)$, is able to accurately regress the mean $\mu(x) = \sin(x)$ for $x \in [-3, 3]$. However, this model fails to capture any notion of uncertainty.





Instead, the DNN f_{θ} can be used to specify a Gaussian model $p(y|x; \theta) = \mathcal{N}(y; \mu_{\theta}(x), \sigma_{\theta}^2(x))$, trained by minimizing the NLL $\mathcal{L}(\theta)$. The model closely matches the true p(y|x) for $x \in [-3, 3]$, accounting for *aleatoric* uncertainty.



For inputs |x| > 3 not seen during training, however, the estimated mean $\mu_{\theta}(x)$ deviates significantly from the true $\mu(x) = \sin(x)$, while the estimated uncertainty $\sigma_{\theta}^2(x)$ remains very small. That is, the model becomes **overconfident** for inputs |x| > 3.



The Gaussian DNN model $p(y|x;\theta) = \mathcal{N}(y;\mu_{\theta}(x),\sigma_{\theta}^2(x))$ can instead be estimated via approximate Bayesian inference, with $M = 1\,000$ samples $\{\theta^{(m)}\}_{m=1}^{M}$ obtained via HMC used in (1), in order to account for both *aleatoric* and *epistemic* uncertainty.



The model now predicts a more reasonable uncertainty $\sigma_{\theta}^2(x)$ in the region with no available training data.

While the estimated mean $\mu_{\theta}(x)$ still deviates from the true $\mu(x) = \sin(x)$ for |x| > 3, the uncertainty $\sigma_{\theta}^2(x)$ also increases accordingly – the model does *not* become overconfident.



While HMC (Hamiltonian Monte Carlo) is considered a "gold standard" method for approximate Bayesian inference, it does not scale well to the large DNNs f_{θ} used in real-world applications.

In practice, among scalable alternatives, it has been shown difficult to beat the simple approach of ensembling. This entails training M identical DNNs by repeatedly minimizing the negative log-likelihood $\mathcal{L}(\theta) = \sum_{i=1}^{N} -\log p(y_i|x_i;\theta)$ with random initialization.

This gives M point estimates $\{\hat{\theta}^{(m)}\}_{m=1}^{M}$ of the DNN model parameters, which can be used as approximate samples for the predictive posterior distribution,

$$p(y|x, \mathcal{D}) = \int p(y|x; \theta) p(\theta|\mathcal{D}) d\theta \approx \frac{1}{M} \sum_{m=1}^{M} p(y|x; \hat{\theta}^{(m)}).$$



In the illustrative 1D regression example, ensembling provides a good approximation of HMC, even for relatively small values of M.



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Given a desired miscoverage rate $\alpha \in]0,1[$, a **prediction interval** $C_{\alpha}(x^*) = [L_{\alpha}(x^*), U_{\alpha}(x^*)] \subseteq \mathbb{R}$ is a function that maps the input x^* onto an interval that should cover the true regression target y^* with probability $1 - \alpha$.

For any set $\{(x_i^*, y_i^*)\}_{i=1}^{N^*}$ of N^* examples, the empirical **interval coverage** is the proportion of inputs for which the prediction interval covers the target,

$$\operatorname{Coverage}(C_{\alpha}) = \frac{1}{N^{\star}} \sum_{i=1}^{N^{\star}} \mathbb{I}\{y_i^{\star} \in C_{\alpha}(x_i^{\star})\}.$$
(2)

If the coverage equals $1 - \alpha$, we say that the prediction intervals are **perfectly** calibrated. Unless stated otherwise, we here set $\alpha = 0.1$. The prediction intervals should thus obtain a coverage of 90%.



Prediction interval: $C_{\alpha}(x^{\star}) = [L_{\alpha}(x^{\star}), U_{\alpha}(x^{\star})] \subseteq \mathbb{R}.$

Empirical interval coverage: Coverage $(C_{\alpha}) = \frac{1}{N^{\star}} \sum_{i=1}^{N^{\star}} \mathbb{I}\{y_i^{\star} \in C_{\alpha}(x_i^{\star})\}.$

With a trained **Gaussian DNN model** $p(y|x;\theta) = \mathcal{N}(y;\mu_{\theta}(x), \sigma_{\theta}^2(x))$, a prediction interval $C_{\alpha}(x^*)$ for a given input x^* can be constructed as,

$$C_{\alpha}(x^{*}) = [\mu_{\theta}(x^{*}) - \sigma_{\theta}(x^{*})\Phi^{-1}(1 - \alpha/2), \ \mu_{\theta}(x^{*}) + \sigma_{\theta}(x^{*})\Phi^{-1}(1 - \alpha/2)],$$

where $\boldsymbol{\Phi}$ is the CDF of the standard normal distribution.



With a trained Gaussian DNN model $p(y|x;\theta) = \mathcal{N}(y;\mu_{\theta}(x),\sigma_{\theta}^2(x))$, a prediction interval $C_{\alpha}(x^*)$ for a given input x^* can be constructed as,

$$C_{\alpha}(x^{\star}) = [\mu_{\theta}(x^{\star}) - \sigma_{\theta}(x^{\star})\Phi^{-1}(1 - \alpha/2), \ \mu_{\theta}(x^{\star}) + \sigma_{\theta}(x^{\star})\Phi^{-1}(1 - \alpha/2)], \ (3)$$

where Φ is the CDF of the standard normal distribution.

With a trained **ensemble** $\{f_{\theta^{(m)}}\}_{m=1}^{M}$ of M such Gaussian DNN models, a single mean $\hat{\mu}$ and variance $\hat{\sigma}^2$ can be computed as,

$$\hat{\mu}(x^{\star}) = \frac{1}{M} \sum_{m=1}^{M} \mu_{\theta^{(m)}}(x^{\star}), \qquad \hat{\sigma}^{2}(x^{\star}) = \frac{1}{M} \sum_{m=1}^{M} \left(\left(\hat{\mu}(x^{\star}) - \mu_{\theta^{(m)}}(x^{\star}) \right)^{2} + \sigma_{\theta^{(m)}}^{2}(x^{\star}) \right),$$

and then plugged into (3) to construct a prediction interval $C_{\alpha}(x^{*})$ for the input x^{*} .

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The general idea of selective prediction is to give a model the option to abstain from outputting predictions for some inputs.

This is achieved by combining the prediction model f_{θ} with an uncertainty function $\kappa_f : \mathcal{X} \to \mathbb{R}$. Given an input x^* , the prediction $f_{\theta}(x^*)$ is output if the uncertainty $\kappa_f(x^*) \leq \tau$, otherwise x^* is rejected and no prediction is made.

The prediction rate is the proportion of inputs for which a prediction is output,

Predition Rate =
$$\frac{1}{N^{\star}} \sum_{i=1}^{N^{\star}} \mathbb{I}\{\kappa_f(x_i^{\star}) \leq \tau\}.$$
 (4)



We combine selective prediction with standard regression methods. A prediction interval $C_{\alpha}(x^{*})$ and predicted target $\hat{y}(x^{*})$ are thus output if and only if $\kappa_{f}(x^{*}) \leq \tau$. Our aim is for this to improve the calibration of the output prediction intervals.

For $\kappa_f(x)$, the variance $\hat{\sigma}^2(x)$ of a Gaussian ensemble could be used, for example.

One could also use some of the various uncertainty scores employed in the rich **out-of-distribution (OOD) detection** literature. In OOD detection, the task is to distinguish in-distribution inputs x, inputs which are similar to those of the training set $\{(x_i, y_i)\}_{i=1}^N$, from out-of-distribution inputs.

A principled approach to OOD detection would be to fit a model of p(x) on the training set. Inputs x for which p(x) is small are then deemed OOD. One can also fit a simple model to the feature vectors g(x), modelling p(x) indirectly.

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6. Motivating Example





Much of the work in the *How Reliable is Your Regression Model's Uncertainty Under Real-World Distribution Shifts*? paper was inspired and motivated by concurrent work on **ECG-based electrolyte prediction**.

Abnormal potassium (K^+) concentration levels in the human body can lead to serious heart conditions. If the concentration could be accurately monitored using an ECG-based regression model, potentially life-threatening conditions could be avoided.
6. Motivating Example





We recently trained a DNN on this task and obtained reasonable regression accuracy (to train the model, we utilized a large-scale dataset of over 290 000 ECGs from adult patients attending emergency departments at Swedish hospitals).

During this work, we started thinking more carefully about the question: **Would it be possible to actually deploy this model in clinical practice at the university hospital?** What requirements would such real-world deployment within a safety-critical domain put on this deep regression model?



The model must at least be **well calibrated**. If it outputs a prediction and a 90% prediction interval for each input, 90% interval coverage should actually be achieved.

Otherwise, if the model becomes overconfident and outputs highly confident yet incorrect predictions, providing uncertainty estimates might just instill a false sense of security – arguably making the model even less suitable for safety-critical deployment.

Moreover, the model must remain well calibrated also under the wide variety of **distribution shifts** that might be encountered during practical deployment.

For example, a model trained on data collected solely at a large urban hospital in the year 2020, for instance, should output well-calibrated predictions also in 2023, for patients both from urban and rural areas.



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7. Proposed Benchmark - Datasets





We collect 8 publicly available datasets for different image-based regression tasks, with various types of distribution shifts (*e.g.*, *train on satellite images captured in densely populated American cities – test on images captured in a rural European area*).

2 synthetic datasets, 6 real-world datasets. 6592 - 20614 training images.



We evaluate regression uncertainty estimation methods mainly in terms of **prediction** interval coverage, $Coverage(C_{\alpha}) = \frac{1}{N^*} \sum_{i=1}^{N^*} \mathbb{I}\{y_i^* \in C_{\alpha}(x_i^*)\}.$

If a method outputs a prediction $\hat{y}(x)$ and a 90% prediction interval $C_{0.1}(x)$ for each input x, does the method actually achieve 90% coverage on the *test* set? I.e., are the prediction intervals calibrated?

We also evaluate in terms of *average interval length* on the val set. This is a natural secondary metric, since a method that achieves a coverage close to $1 - \alpha$ but outputs extremely large intervals for all inputs x, would not be particularly useful in practice.



For methods based on selective prediction, the only difference is that predictions $\hat{y}(x)$ and prediction intervals $C_{\alpha}(x)$ are output only for some test inputs x (iff $\kappa_f(x) \leq \tau$). The prediction interval coverage is thus computed only on this subset of test.

For these methods, the proportion of inputs for which a prediction actually is output is another natural secondary metric. We thus also evaluate in terms of the *prediction rate* $\frac{1}{N^*} \sum_{i=1}^{N^*} \mathbb{I}\{\kappa_f(x_i^*) \leq \tau\}$ on test.

If a coverage close to $1 - \alpha$ is achieved with a very low prediction rate, the method might still not be practically useful in certain applications.



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We evaluate five common regression uncertainty estimation methods, which all output a 90% prediction interval $C_{0.1}(x)$ and a predicted target $\hat{y}(x) \in C_{0.1}(x)$ for each input.

Two of these methods we also combine with selective prediction, utilizing four different uncertainty functions $\kappa_f(x)$.

In total, we evaluate 10 different methods.

We calibrate the prediction intervals, for each of the 10 methods, such that exactly 90% interval coverage is obtained on the val set. Ideally, the coverage should then not change from the val set to the test set.



Conformal Prediction.

Ensemble.

Gaussian.

Gaussian Ensemble.

Quantile Regression.



Gaussian + Selective GMM.

Combining a Gaussian model p(y|x; θ) = N(y; μ_θ(x), σ²_θ(x)) with selective prediction. A GMM is fit to the feature vectors {g(x_i)}^N_{i=1} of the training set. The GMM likelihood is then taken as the uncertainty score, κ_f(x) = -GMM(g(x)).

Gaussian + Selective kNN.

 The average distance from g(x) to its k nearest neighbors among the train feature vectors {g(x_i)}^N_{i=1} is taken as the uncertainty score, κ_f(x) = kNN(g(x)).

Gaussian + Selective Variance.

Gaussian Ensemble + Selective GMM.

Gaussian Ensemble + Selective Ensemble Variance



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(1/4) All methods are well calibrated on baseline datasets with no distribution shift, but become highly overconfident in many realistic scenarios. Uncertainty estimation methods must therefore be evaluated using sufficiently challenging benchmarks. Otherwise, one might be lead to believe that methods will be more reliable during real-world deployment than they actually are.

(2/4) Conformal prediction methods have commonly promoted theoretical coverage guarantees, but these depend on an assumption that is unlikely to hold in many practical applications. Consequently, also these methods can become highly overconfident in realistic scenarios. If the underlying assumptions are not examined critically by practitioners, such theoretical guarantees risk instilling a false sense of security – making these models even less suitable for safety-critical deployment.



(3/4) The clear performance difference between synthetic and real-world datasets observed for selective prediction methods based on feature-space density is a very interesting direction for future work. If the reasons for this performance gap can be understood, an uncertainty estimation method that stays well calibrated across all datasets could potentially be developed.

(4/4) Selective prediction methods based on feature-space density perform well relative to other methods (as expected based on their state-of-the-art OOD detection performance), but are also overconfident in many cases. Only comparing the relative performance of different methods is therefore not sufficient. To track if actual progress is being made towards the ultimate goal of truly reliable uncertainty estimation methods, benchmarks must also evaluate method performance in an absolute sense.



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